DSC291: Machine Learning with Few Labels

Variational Inference

Zhiting Hu Lecture 7, January 24, 2023



Recap: EM Algorithm

• The EM algorithm is coordinate-decent on $F(q, \theta)$

$$\circ$$
 E-step: $q^{t+1} = \arg\min_{q} F\left(q, \theta^{t}\right) = p(\mathbf{z}|\mathbf{x}, \theta^{t})$

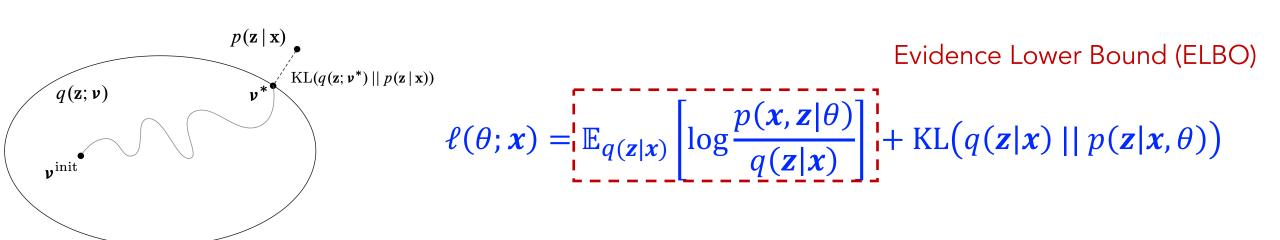
$$\circ \quad \text{M-step:} \quad \theta^{t+1} = \arg\min_{\theta} F\left(q^{t+1}, \theta^{t}\right) = \operatorname{argmax}_{\theta} \sum_{\mathbf{z}} q^{t+1}(\mathbf{z}|\mathbf{x}) \log p(\mathbf{x}, \mathbf{z}|\theta)$$

$$\ell(\theta; \mathbf{x}) = \mathbb{E}_{q(\mathbf{z}|\mathbf{x})} \left[\log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{q(\mathbf{z}|\mathbf{x})} \right] + \text{KL}(q(\mathbf{z}|\mathbf{x}) || p(\mathbf{z}|\mathbf{x}, \theta))$$
$$= -F(q, \theta) + \text{KL}(q(\mathbf{z}|\mathbf{x}) || p(\mathbf{z}|\mathbf{x}, \theta))$$

• Limitation: need to be able to compute $p(\mathbf{z}|\mathbf{x},\theta)$, not possible for more complicated models --- solution: Variational inference

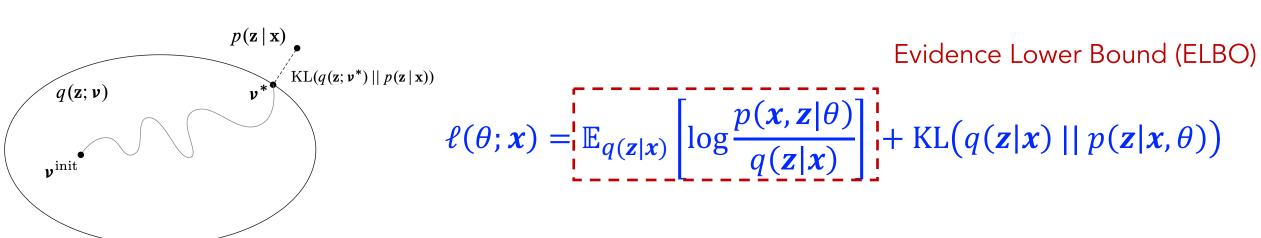
Recap: Variational Inference

- We often cannot compute posteriors $p(\mathbf{z}|\mathbf{x},\theta)$, and so we need to approximate them, using variational methods.
- In variational Bayes, we'd like to find an approximation within some family that minimizes the KL divergence to the posterior, but we can't directly minimize this
- Therefore, we defined the ELBO, which we can maximize, and this is equivalent to minimizing the KL divergence.



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• How do we choose the variational family $q(\mathbf{z}|\mathbf{x}, \mathbf{v})$?

Mean Field Variational Inference

- A popular family of variational approximations
- In this type of variational inference, we assume the variational distribution over the latent variables factorizes as

$$q(\mathbf{z}) = q(z_1, \dots, z_m) = \prod_{j=1}^m q(z_j)$$

- (where we omit variational parameters for ease of notation)
- We refer to $q(z_j)$, the variational approximation for a single latent variable, as a "local variational approximation"
- In the above expression, the variational approximation $q(z_j)$ over each latent variable z_i is independent

Mean Field Variational Inference

• Note that this is a fairly general setup; we can also partition the latent variables z_1, \ldots, z_m into R groups z_{G_1}, \ldots, z_{G_R} , and use the approximation:

$$q(z_1,\ldots,z_m)=q(z_{G_1},\ldots,z_{G_R})=\prod_{r=1}^R q(z_{G_r})$$

Often called "generalized mean field" versus (the above) "naïve mean field".

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- Typically, this approximation does not contain the true posterior (because the latent variables are dependent).
 - E.g.: in the (Bayesian) mixture of Gaussians model, all of the cluster assignments z_i for $i=1,\ldots,n$ are dependent on each other and on the cluster locations $\mu_{1:K}$ given data.

How do we optimize the ELBO in mean field variational inference?

- Typically, we use coordinate ascent optimization.
- I.e. we optimize each latent variable's variational approximation $q(z_j)$ in turn while holding the others fixed.
 - At each iteration we get an updated "local" variational approximation.
 - And we iterate through each latent variable until convergence.

Recall that the ELBO is defined as:

$$\mathcal{L} = \mathbb{E}_q[\log p(\mathbf{x}, \mathbf{z})] - \mathbb{E}_q[\log q(\mathbf{z})]$$

 Note that we can decompose the entropy term of the ELBO (using the mean field variational approximation) as:

$$\mathbb{E}_q\left[\log q(z_{1:m})
ight] = \sum_{j=1}^m \mathbb{E}_{q_j}\left[\log q(z_j)
ight]$$

• Therefore, under the mean field approximation, the ELBO can be written:

$$\mathcal{L} = \mathbb{E}_{q_j} \mathbb{E}_{q_{-j}} [\log p(\mathbf{x}, \mathbf{z})] - \sum_{j=1}^{m} \mathbb{E}_{q_j} [\log q(z_j)]$$

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- Next, we want to derive the coordinate ascent update for a latent variable z_i , keeping all other latent variables fixed.
 - o i.e. we want the $\operatorname{argmax}_{q_i} \mathcal{L}$.
- Removing the parts that do not depend on $q(z_i)$, we can write:

$$\mathcal{L} = \mathbb{E}_{q_i} \mathbb{E}_{q_{-i}} [\log p(\mathbf{x}, \mathbf{z})] - \mathbb{E}_{q_i} [\log q(z_i)] + \text{const.}$$

• To find this argmax, we take the derivative of \mathcal{L} w.r.t $q(z_j)$ and and set the derivative to zero :

$$\frac{d\mathcal{L}}{dq(z_i)} = \mathbb{E}_{q_j} \mathbb{E}_{q_{-j}} [\log p(\mathbf{x}, \mathbf{z})] - \log q(z_j) - 1 = 0$$

• From this, we arrive at the coordinate ascent update:

$$q^*(z_j) \propto \exp \left\{ \mathbb{E}_{q_{-j}}[\log p(\mathbf{x}, \mathbf{z})] \right\}$$

The coordinate ascent update:

$$q^*(z_j) \propto \exp \left\{ \mathbb{E}_{q_{-j}}[\log p(\mathbf{x}, \mathbf{z})] \right\}$$

- The optimal solution for factor $q(z_j)$ is obtained simply by considering the log of the joint distribution over all observed and latent variables and then taking the expectation with respect to all of the other factors $q(z_k)$, $k \neq j$, then taking exponential and normalizing
- Note that the only assumption we made so far is the mean-field factorization: $q(\mathbf{z}) = q(z_1, \dots, z_m) = \prod_{i=1}^m q(z_j)$

• We haven't yet made any assumptions on the form of $q(z_i)$

Simple example:

• Consider a univariate Gaussian distribution $p(x) = \mathcal{N}(x|\mu,\tau^{-2})$, given a dataset $\mathcal{D} = \{x_1, ..., x_N\}$:

$$p(\mathcal{D}|\mu,\tau) = \left(\frac{\tau}{2\pi}\right)^{N/2} \exp\left\{-\frac{\tau}{2} \sum_{n=1}^{N} (x_n - \mu)^2\right\}$$
$$p(\mu|\tau) = \mathcal{N}\left(\mu|\mu_0, (\lambda_0 \tau)^{-1}\right)$$
$$p(\tau) = \operatorname{Gam}(\tau|a_0, b_0)$$

- o $\operatorname{Gam}(\tau|a_0,b_0)=\frac{1}{\Gamma(a)}b^a\lambda^{a-1}\exp(-b\lambda)$: gamma distribution
- For this simple problem the posterior distribution can be found exactly. But we use it as an example for tutorial anyway

$$q^*(z_j) \propto \exp \left\{ \mathbb{E}_{q_{-j}}[\log p(\mathbf{x}, \mathbf{z})] \right\}$$

Simple example:

$$p(\mathcal{D}|\mu,\tau) = \left(\frac{\tau}{2\pi}\right)^{N/2} \exp\left\{-\frac{\tau}{2} \sum_{n=1}^{N} (x_n - \mu)^2\right\} \qquad p(\mu|\tau) = \mathcal{N}\left(\mu|\mu_0, (\lambda_0 \tau)^{-1}\right) \\ p(\tau) = \operatorname{Gam}(\tau|a_0, b_0)$$

- Introduce the factorized variational approximation: $q(\mu, \tau) = q_{\mu}(\mu)q_{\tau}(\tau)$
- Solution to q_{μ} :

$$\ln q_{\mu}^{\star}(\mu) = \mathbb{E}_{\tau} \left[\ln p(\mathcal{D}|\mu,\tau) + \ln p(\mu|\tau) \right] + \text{const}$$

$$= -\frac{\mathbb{E}[\tau]}{2} \left\{ \lambda_0 (\mu - \mu_0)^2 + \sum_{n=1}^{N} (x_n - \mu)^2 \right\} + \text{const.}$$

• We can see q_{μ}^* is a Gaussian $\mathcal{N}(x|\mu_N,\lambda_N^{-1})$:

$$\mu_N = \frac{\lambda_0 \mu_0 + N\overline{x}}{\lambda_0 + N}$$
 $\lambda_N = (\lambda_0 + N)\mathbb{E}[\tau]$

$$q^*(z_j) \propto \exp \left\{ \mathbb{E}_{q_{-j}}[\log p(\mathbf{x}, \mathbf{z})] \right\}$$

Simple example:

$$p(\mathcal{D}|\mu,\tau) = \left(\frac{\tau}{2\pi}\right)^{N/2} \exp\left\{-\frac{\tau}{2} \sum_{n=1}^{N} (x_n - \mu)^2\right\} \qquad p(\mu|\tau) = \mathcal{N}\left(\mu|\mu_0, (\lambda_0 \tau)^{-1}\right) \\ p(\tau) = \operatorname{Gam}(\tau|a_0, b_0)$$

- Introduce the factorized variational approximation: $q(\mu, \tau) = q_{\mu}(\mu)q_{\tau}(\tau)$
- Solution to q_{τ} : $\ln q_{\tau}^{\star}(\tau) = \mathbb{E}_{\mu} \left[\ln p(\mathcal{D}|\mu,\tau) + \ln p(\mu|\tau) \right] + \ln p(\tau) + \text{const}$ $= (a_0 - 1) \ln \tau - b_0 \tau + \frac{N}{2} \ln \tau$ $-\frac{\tau}{2} \mathbb{E}_{\mu} \left[\sum_{n=1}^{N} (x_n - \mu)^2 + \lambda_0 (\mu - \mu_0)^2 \right] + \text{const}$

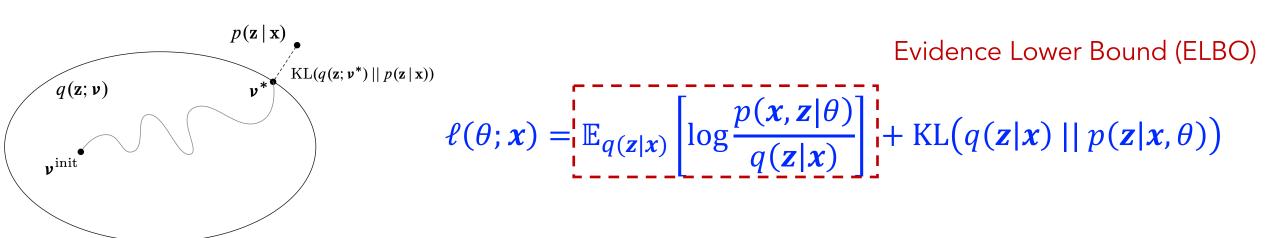
• We can see q_{τ}^* is a gamma distribution $Gam(\tau|a_N,b_N)$:

$$a_N = a_0 + \frac{N}{2}$$

$$b_N = b_0 + \frac{1}{2} \mathbb{E}_{\mu} \left[\sum_{n=1}^{N} (x_n - \mu)^2 + \lambda_0 (\mu - \mu_0)^2 \right]$$

Quick Recap

- We often cannot compute posteriors, and so we need to approximate them, using variational methods.
- In variational Bayes, we'd like to find an approximation within some family that minimizes the KL divergence to the posterior, but we can't directly minimize this
- Therefore, we defined the ELBO, which we can maximize, and this is equivalent to minimizing the KL divergence.



Quick Recap

• We defined a family of approximations called "mean field" approximations, in which there are no dependencies between latent variables $\frac{m}{}$

$$q(\mathbf{z}) = q(z_1, \dots, z_m) = \prod_{j=1} q(z_j)$$

 We optimize the ELBO with coordinate ascent updates to iteratively optimize each local variational approximation under mean field assumptions

$$q^*(z_j) \propto \exp \left\{ \mathbb{E}_{q_{-j}}[\log p(\mathbf{x}, \mathbf{z})] \right\}$$

Key Takeaways

KL Divergence

- $KL(q(\mathbf{x}) \mid\mid p(\mathbf{x})) = \sum_{\mathbf{x}} q(\mathbf{x}) \log \frac{q(\mathbf{x})}{p(\mathbf{x})}$
- The key equation of EM and VI

Evidence Lower Bound (ELBO)

$$\ell(\theta; \mathbf{x}) = \mathbb{E}_{q(\mathbf{z}|\mathbf{x})} \left[\log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{q(\mathbf{z}|\mathbf{x})} \right] + \text{KL} \left(q(\mathbf{z}|\mathbf{x}) \mid\mid p(\mathbf{z}|\mathbf{x}, \theta) \right)$$

- Free energy $F(q, \theta)$
- EM: E-step and M-step optimizing ELBO w.r.t q and θ
- Mean-field VI: optimizing ELBO w.r.t factorized q components

Stochastic VI; Black-box VI

VI with coordinate ascent

Example: Bayesian mixture of Gaussians

• Treat the mean μ_k and cluster proportion π as latent variables

$$\mu_k \sim \mathcal{N}(0, \tau^2) \text{ for } k = 1, \dots, K$$

$$\pi \sim Dirichlet(\boldsymbol{\alpha})$$

• For each data i = 1, ..., n

$$z_i \sim \operatorname{Cat}(\pi).$$

 $x_i \sim \mathcal{N}(\mu_{z_i}, \sigma^2).$

- We have
 - observed variables $x_{1:n}$
 - o latent variables $\mu_{1:k}$, π and $z_{1:n}$
 - Hyper-parameters $\{\tau^2, \sigma^2\}$

VI with coordinate ascent

Example: Bayesian mixture of Gaussians Assume mean-field $q(\mu_{1:K}, \pi, z_{1:n}) = \prod_k q(\mu_k) q(\pi) \prod_i q(z_i)$

- Initialize the global variational distributions $q(\mu_k)$ and $q(\pi)$
- Repeat:
 - For each data example $i \in \{1,2,...,D\}$
 - Update the local variational distribution $q(z_i)$
 - End for
 - Update the global variational distributions $q(\mu_k)$ and $q(\pi)$
- Until ELBO converges

• What if we have millions of data examples? This could be very slow.

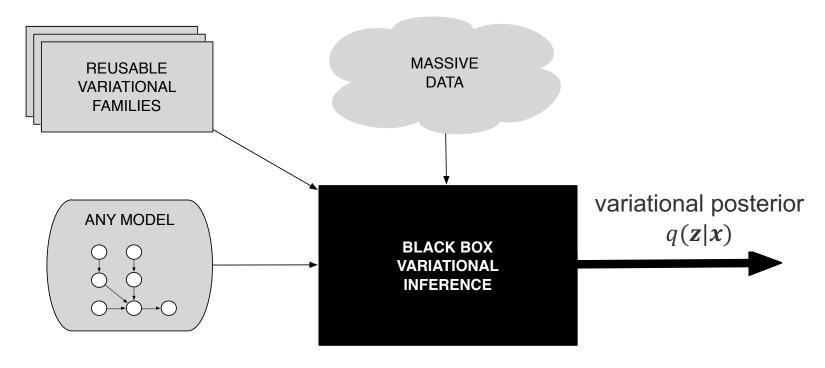
Stochastic VI

Example: Bayesian mixture of Gaussians

Assume mean-field $q(\mu_{1:K}, \pi, z_{1:n}) = \prod_k q(\mu_k) q(\pi) \prod_i q(z_i)$

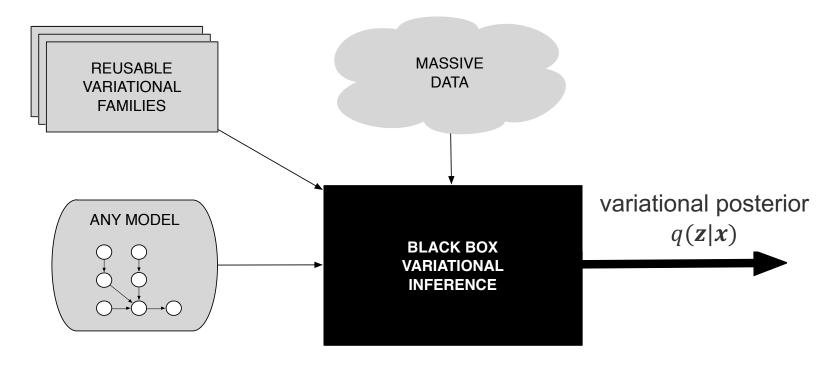
- ullet Initialize the global variational distributions $q(\mu_k)$ and $q(\pi)$
- Repeat:
 - Sample a data example $i \in \{1,2,...,D\}$
 - Update the local variational distribution $q(z_i)$
 - Update the global variational distributions $q(\mu_k)$ and $q(\pi)$ with natural gradient ascent
- Until ELBO converges
- (Setting natural gradient = 0 gives the traditional mean-field update)

- We have derived variational inference specific for Bayesian Gaussian (mixture) models
- There are innumerable models
- Can we have a solution that does not entail model-specific work?



- Easily use variational inference with any model
- Perform inference with massive data
- No mathematical work beyond specifying the model

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- Sample from q(.)
- Form noisy gradients (without model-specific computation)
- Use stochastic optimization

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- Probabilistic model: x -- observed variables, z -- latent variables
- Variational distribution $q_{\lambda}(\mathbf{z}|\mathbf{x})$ with parameters λ , e.g.,
 - Gaussian mixture distribution:
 - "A Gaussian mixture model is a universal approximator of densities, in the sense that any smooth density can be approximated with any specific nonzero amount of error by a Gaussian mixture model with enough components." (Deep Learning book, pp.65)
 - Deep neural networks
- ELBO:

$$\mathcal{L}(\lambda) = \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log p(\mathbf{x}, \mathbf{z})] - \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log q(\mathbf{z}|\lambda)]$$

• Want to compute the gradient w.r.t variational parameters λ

The General Problem: Computing Gradients of Expectations

• When the objective function \mathcal{L} is defined as an expectation of a (differentiable) test function $f_{\lambda}(\mathbf{z})$ w.r.t. a probability distribution $q_{\lambda}(\mathbf{z})$

$$\mathcal{L} = \mathbb{E}_{q_{\lambda}(\mathbf{z})}[f_{\lambda}(\mathbf{z})]$$

- Computing exact gradients w.r.t. the parameters λ is often unfeasible
- Need stochastic gradient estimates
 - The score function estimator (a.k.a log-derivative trick, REINFORCE)
 - The reparameterization trick (a.k.a the pathwise gradient estimator)

Computing Gradients of Expectations w/ score function

- Loss: $\mathcal{L} = \mathbb{E}_{q_{\lambda}(\mathbf{z})}[f_{\lambda}(\mathbf{z})]$
- Log-derivative trick: $\nabla_{\lambda} q_{\lambda} = q_{\lambda} \nabla_{\lambda} \log q_{\lambda}$
- Gradient w.r.t. λ:

$$\nabla_{\lambda} \mathcal{L} = \mathbb{E}_{q_{\lambda}(\mathbf{z})} [f_{\lambda}(\mathbf{z}) \nabla_{\lambda} \log q_{\lambda}(\mathbf{z}) + \nabla_{\lambda} f_{\lambda}(\mathbf{z})]$$

- o score function: the gradient of the log of a probability distribution
- ullet Compute noisy unbiased gradients with Monte Carlo samples from q_λ

$$\nabla_{\lambda} \mathcal{L} \approx \frac{1}{S} \sum_{s=1}^{S} f_{\lambda}(\mathbf{z}_{s}) \nabla_{\lambda} \log q_{\lambda}(\mathbf{z}_{s}) + \nabla_{\lambda} f_{\lambda}(\mathbf{z}_{s})$$
 where $\mathbf{z}_{s} \sim q_{\lambda}(\mathbf{z})$

- Pros: generally applicable to any distribution $q(z|\lambda)$
- Cons: empirically has high variance → slow convergence
 - To reduce variance: Rao-Blackwellization, control variates, importance sampling, ...

Computing Gradients of Expectations w/ reparametrization trick

- Loss: $\mathcal{L} = \mathbb{E}_{q_{\lambda}(\mathbf{z})}[f_{\lambda}(\mathbf{z})]$
- Assume that we can express the distribution $q_{\lambda}(z)$ with a transformation

$$\begin{array}{l}
\epsilon \sim s(\epsilon) \\
z = t(\epsilon, \lambda)
\end{array} \iff z \sim q(z|\lambda)$$

E.g.,

$$\begin{array}{l} \epsilon \sim Normal(0,1) \\ z = \epsilon \sigma + \mu \end{array} \iff z \sim Normal(\mu, \sigma^2)$$

Reparameterization gradient

$$\mathcal{L} = \mathbb{E}_{\epsilon \sim S(\epsilon)}[f_{\lambda}(\mathbf{z}(\epsilon, \lambda))]$$

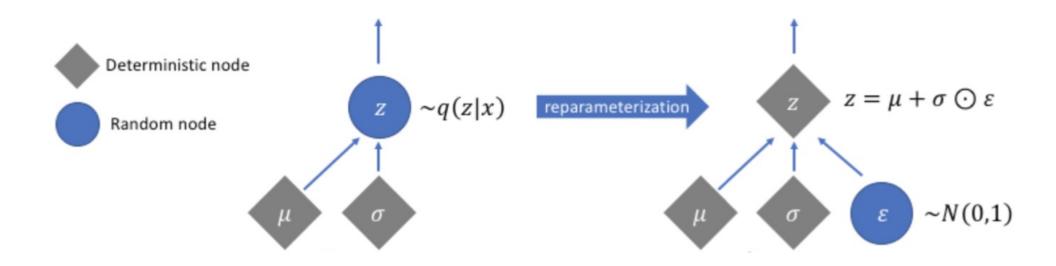
$$\nabla_{\lambda} \mathcal{L} = \mathbb{E}_{\epsilon \sim S(\epsilon)}[\nabla_{\mathbf{z}} f_{\lambda}(\mathbf{z}) \nabla_{\lambda} t(\epsilon, \lambda)]$$

- Pros: empirically, lower variance of the gradient estimate
- Cons: Not all distributions can be reparameterized

Reparameterization trick

• Reparametrizing Gaussian distribution

$$\begin{array}{l} \epsilon \sim Normal(0,1) \\ z = \epsilon \sigma + \mu \end{array} \iff z \sim Normal(\mu, \sigma^2)$$



[Courtesy: Tansey, 2016]

Reparameterization trick

• Reparametrizing Gaussian distribution

$$\begin{array}{l} \epsilon \sim Normal(0,1) \\ z = \epsilon \sigma + \mu \end{array} \iff z \sim Normal(\mu, \sigma^2)$$

- Other reparameterizable distributions: $\epsilon \sim Uniform(\epsilon)$ • Tractable inverse CDF F^{-1} : $z = F^{-1}(\epsilon)$ $\Leftrightarrow z \sim q(z)$
 - Exponential, Cauchy, Logistic, Rayleigh, Pareto, Weibull, Reciprocal, Gompertz, Gumbel, Erlang
 - Location-scale:
 - Laplace, Elliptical, Student's t, Logistic, Uniform, Triangular, Gaussian
 - Composition:
 - Log-Normal (exponentiated normal) Gamma (sum of exponentials) Dirichlet (sum of Gammas) Beta, Chi-Squared, F

[Courtesy: Tansey, 2016]

Computing Gradients of Expectations: Summary

- Loss: $\mathcal{L} = \mathbb{E}_{q_{\lambda}(\mathbf{z})}[f_{\lambda}(\mathbf{z})]$
- Score gradient

$$\nabla_{\lambda} \mathcal{L} = \mathbb{E}_{q_{\lambda}(\mathbf{z})} [f_{\lambda}(\mathbf{z}) \nabla_{\lambda} \log q_{\lambda}(\mathbf{z}) + \nabla_{\lambda} f_{\lambda}(\mathbf{z})]$$

- Pros: generally applicable to any distribution $q(z|\lambda)$
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- Reparameterization gradient

$$\nabla_{\lambda} \mathcal{L} = \mathbb{E}_{\epsilon \sim s(\epsilon)} [\nabla_{\mathbf{z}} f_{\lambda}(\mathbf{z}) \nabla_{\lambda} t(\epsilon, \lambda)]$$

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- Variational distribution $q_{\lambda}(\mathbf{z}|\mathbf{x})$ with parameters λ , e.g.,
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 - Deep neural networks $\mathcal{L}(\lambda) \triangleq \mathrm{E}_{q_{\lambda}(z)}[\log p(x,z) \log q(z)].$
- ELBO:

$$\mathcal{L}(\lambda) = \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log p(\mathbf{x}, \mathbf{z})] - \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log q(\mathbf{z}|\lambda)]$$

• Want to compute the gradient w.r.t variational parameters λ

BBVI with the score gradient

• ELBO:

$$\mathcal{L}(\lambda) = \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log p(\mathbf{x}, \mathbf{z})] - \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log q(\mathbf{z}|\lambda)]$$

• Gradient w.r.t. λ (using the log-derivative trick)

$$\nabla_{\lambda} \mathcal{L} = \mathrm{E}_{q}[\nabla_{\lambda} \log q(z|\lambda)(\log p(x,z) - \log q(z|\lambda))]$$

 Compute noisy unbiased gradients of the ELBO with Monte Carlo samples from the variational distribution

$$abla_{\lambda} \mathcal{L} pprox rac{1}{S} \sum_{s=1}^{S}
abla_{\lambda} \log q(z_s | \lambda) (\log p(x, z_s) - \log q(z_s | \lambda)),$$
where $z_s \sim q(z | \lambda)$.

BBVI with the reparameterization gradient

• ELBO:

$$\mathcal{L}(\lambda) = \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log p(\mathbf{x}, \mathbf{z})] - \mathbb{E}_{q(\mathbf{z}|\lambda)}[\log q(\mathbf{z}|\lambda)]$$

Gradient w.r.t. λ

$$\begin{array}{l} \epsilon \sim s(\epsilon) \\ z = t(\epsilon, \lambda) \end{array} \iff z \sim q(z|\lambda)$$

$$\nabla_{\lambda} \mathcal{L} = \mathbb{E}_{\epsilon \sim s(\epsilon)} \left[\nabla_z [\log p(x, z) - \log q(z)] \nabla_{\lambda} t(\epsilon, \lambda) \right]$$

Questions?

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